Using Network Data to Measure Social Returns and Improve Targeting of Crime-Reduction Interventions

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All opinions are those of the authors and do not represent the position of any government organization or data provider.

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The high cost of crime and violence

- ► Acts of violence kill over 160 per day in the U.S. (CDC 2019)
 - And 2.3 million people live behind bars (Prison Policy Initiative 2020)
- Stark inequalities in who bears these costs
 - ▶ Violent-crime arrests 5 Xs higher for Black than White young people (OJJDP 2014)
 - ▶ With lifelong consequences on social & economic outcomes (Aizer & Doyle 2015, Mueller-Smith 2015, Nelson & Sheridan 2011)
- Social costs of gun violence alone at least \$100 billion/year, maybe > \$500 billion (Cook & Ludwig 2000, Gobbo 2023)

Randomized controlled trials (RCTs) and crime prevention

- ▶ Hundreds of RCTs over 4 decades have tested ways to address these problems
 - Policing strategies (e.g., Braga et al. 1999, 2006, 2017; Owens et al. 2018; Sherman & Berk 1984; Sherman & Weisburd 1995)
 - ► Employment & re-entry (e.g., Cook et al. 2015; MDRC 1980; Kemper et al. 1981; Valentine & Redcross 2015),
 - ▶ Urban environment (e.g., Branas et al. 2018; Chalfin et al. 2021)
 - ► Education and skill development (e.g., Armstrong et al. 2003; Dodge et al 2007; Schweinhart et al. 2007)
- ▶ Influential b/c convincingly isolate causal effect of treatment vs control condition
- Most assume one individual's behavior doesn't affect others (SUTVA)
 - Also true of many non-RCT quasi-experiments

In fact, crime is social

- ➤ 50-85% of offenders offend with others, usually as youth when most crime occurs (Conway & McCord 2002; Sarnecki 2001)
- ▶ Peer effects in crime are well established
 - Exogeneous changes in exposure to A changes B's crime (Bayer et al. 2009; Bhuller et al. 2018; Billings et al. 2019; Damm & Gorinas 2020; Drago & Galbiati 2012; Dominguez 2021; Dustmann & Landersø 2021; Norris et al. 2021; Stevenson 2017; Philippe 2017)
- ▶ But existence of peer effects ≠ estimate of how an intervention's crime change spreads
 - ► Requires knowing how many people A affects, if effect varies with who A and B are/how they're connected, if changing A's behavior or just exposure to A matters

Our paper: how do crime changes spread through networks?

- ► Challenge 1: Measure social networks
- ▶ Use administrative data on \sim 2m people \rightarrow Chicago Police Department (arrests & reported victimizations 2005-21) and Chicago Public Schools records (2009-20)
- ► Co-arrest, co-victimization, same classes, same residence, geographic proximity
 - ► Clearly misses some strong friendship ties, but non-friend ties may matter (Granovetter 1983, Patacchini & Zenou 2008)
 - ▶ If these ties do matter, makes it feasible to look at crime spillovers retrospectively, at large scale

Our paper: how do crime changes spread through networks?

- **Challenge 2: Overcome classic identification issues** → reflection problem. endogeneity of ties, common shocks (Manski 1993, Angrist 2014)
- Leverage 4 existing, large-scale RCTs of violence-reducing interventions in Chicago Davis & Heller 2020, Heller 2014, Heller et al. 2017, Bhatt et al 2024)
 - Given baseline tie to RCT member, exposure to T is random with known probability
 - ► Aronow & Samii 2017 → probability as IPW, conservative inference
 - Pooling RCTs may help solve challenges in handful of previous efforts to do this
 - Under-powered, potentially confounded (Abdul-Razzak et al, in progress, Dominguez 2023, Wood & Papachristos 2019)

What we do

- 1. Today: Describe networks, estimate preliminary indirect exposure effects
 - Exposure = connected to at least one peer assigned to treatment (ITT)
 - ▶ Within RCT exposure effect: what initial ITTs missed for original samples
 - Out of RCT exposure effect: impact on those not in original RCTs
 - ▶ For now, 1st-degree peer + any exposure, no geography or overlapping networks yet
- 2. Future: Estimate counterfactual targeting mechanisms \rightarrow optimal targeting
 - Use heterogeneity to tease out behavioral mechanisms
 - ▶ Build model incorporating effects on future tie formation, network position

Intervention spillovers and social networks

- Multiple literatures developed methods to analyze peer effects, network diffusion
 - In development: how innovations diffuse through social networks, change network structure (Banerjee et al. 2013, Beaman et al. 2020, Beaman & Magruer 2012, Bhattacharya et al. 2013, Breza & Chandrasekhar 2019, Cai et al. 2015, Comola & Prina 2014, Feigenberg et al. 2013, Miguel & Kremer 2007, Miller & Mobarak 2014, Oster & Thornton 2012)
 - ▶ In education: how exposure to different peers or treating peers matters for learning/behavior (Babcock & Hartman 2010, Chaisemartin & Navarrete 2020, Dinarte & Egana-del Sol 2022, Paluck et al. 2016, Sacerdote 2001, review in Sacerdote 2011)
- Empirical crime literature: more focused on *you* when someone you know/live with gets (un)incarcerated, your school peers or cellmates change, you move, etc.
 - Exception: bullying interventions via reported friend networks (Hu 2024; Paluck, Shepherd & Aronow 2016)

What we are learning

- ► These networks capture many relationships
 - ~550 people exposed for every 100 people treated
- Peers of study members are different
 - Both from the study population and across network types
- Spillovers concentrated among peers also in the original RCTs
 - Observable and unobservable selection into the RCTs
- Spillovers vary by type of network connection
 - ▶ Diffusion among denser networks, substitution among closer relationships
- Spillovers are large
 - Current estimates suggest original RCTs understated violence ↓ by 40-80%

Random variation: 4 violence-reducing RCTs in Chicago

- 1. **Becoming a Man** (BAM 2009-10, n = 2.740)
 - School-based cognitive behavioral therapy (CBT) intervention for 7th-10th grade boys
 - ▶ ITT/LATE: Violent-crime arrests $\downarrow 21/45\%$ in year 1
- 2 & 3. One Summer Chicago Plus (OSC+ 2012 n = 1,634, OSC+ 2013 = 5,216)
 - SYEP+CBT w/ school-based recruiting in 2012, broader reach into legal system in 2013
 - ► Violent-crime arrests ↓ 21-41% / 32-42% in year 1
 - Not just incapacitation, continued decline post-program
- 4. Rapid Employment & Development Initiative (READI 2016-21, n = 2,456)
 - ▶ 18m job, CBT, wraparound services for men (18+) at highest risk of shooting/being shot
 - ► After 20m. 45/65% \downarrow shooting & homicide arrests (adj. p=0.13), 38/48% \downarrow S&H victimization for pre-specified subgroup

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Outcomes

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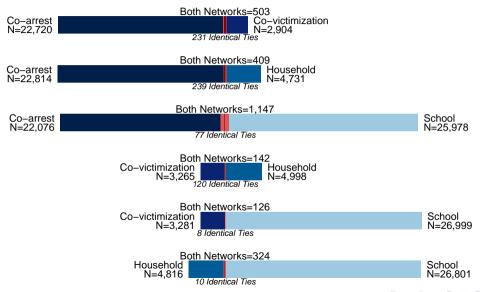
Why a shock to criminal behavior might spill over

- **Diffusion of behavioral change**: information, changes in attitudes/beliefs/time use get transmitted \rightarrow peer crime \downarrow (e.g. Akerlof & Kranton 2010, Pattacchini & Zenou 2009)
- 2. **Key players**: lose key member of criminal team, social influencer \rightarrow peer crime \downarrow (e.g. Lindquist & Zenou 2014, Tankard & Paluck 2016)
- 3. Crime as function of opportunity: if one youth desists, another may take his place \rightarrow peer crime \uparrow (e.g. Cook 1986)
- 4. **Skill complementarity**: co-offending improves productivity \rightarrow peer crime \downarrow , or perhaps substitution to new peer with similar skills \(\frac{1}{2}\) (e.g. Tremblay 1993; Weerman 2003)

Constructing networks

- 1. **Co-arrest**: Arrested together or for same incident in 5 years prior to randomization
 - ▶ 42% in network, 2.8 ties/RCT member (6.8|co arrest > 0)
- 2. **Co-victimization**: Victimized in same incident in 5 years prior to randomization
- ▶ 15% in network, 0.27 ties/RCT member (1.8|co-victim>0)
- 3. Household: Share parent/guardian name, address & last name in CPS data
 - \triangleright 29% in network, 0.44 ties/RCT member (1.5|sibling > 0) Details
 - \triangleright Limit to age > 12 for crime outcomes
- 4. Shared classes: Clustering algorithm \rightarrow many shared courses in prior year
 - Schedule data currently only OSC 1 & 2, BAM coming, READI schooling often old
 - ▶ 48% in network, 6.7 ties/RCT member, (13.9|sharedclass > 0) Details
 - Students in community share avg of 5 of 16 courses, 0.5 outside community
- 5. **Neighborhood** (not yet)
 - ► Using CPS and CPD addresses, live within m mile radius in prior year

Size and overlap of the networks: first-degree peers



RCT study sample descriptive statistics

| | BAM/OSC | READI | | |
|-----------------------|---------|--------|--|--|
| | In-RCT | In-RCT | | |
| N | 9,590 | 2,456 | | |
| Age | 17.4 | 25.7 | | |
| Black | 0.85 | 0.96 | | |
| Male | 0.90 | 1.00 | | |
| Prior Arrests | 2.5 | 16.7 | | |
| Share Co-Arrests | 0.21 | 0.32 | | |
| Degree | 2.23 | 5.02 | | |

Subset in co-arrest network & tied to another study member

| | $_{ m BAM/OSC}$ | | | READI | | |
|------------------|-----------------|------------------------------------|--------|------------------------------|--|--|
| | In-RCT | In-RCT + Potentially Exposed | In-RCT | In-RCT + Potentially Exposed | | |
| N | 9,590 | 1,396 | 2,456 | 975 | | |
| Age | 17.4 | 18.0 | 25.7 | 23.6 | | |
| Black | 0.85 | 0.91 | 0.96 | 1.00 | | |
| Male | 0.90 | 0.98 | 1.00 | 1.00 | | |
| Prior Arrests | 2.5 | 8.8 | 16.7 | 20.3 | | |
| Share Co-Arrests | 0.21 | 0.63 | 0.32 | 0.45 | | |
| Degree | 2.23 | 9.56 | 5.02 | 9.18 | | |

Co-arrest linkages: First degree peers are different

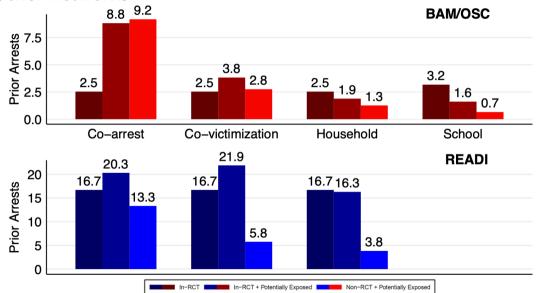
| | BAM/OSC | | | READI | | |
|------------------|---------|------------------------------------|-------------------------------------|------------------------|------------------------------------|-------------------------------------|
| | In-RCT | In-RCT + Potentially Exposed | Non-RCT + Potentially Exposed | In-RCT | In-RCT + Potentially Exposed | Non-RCT + Potentially Exposed |
| N | 9,590 | 1,396 | 14,257 | 2,456 | 975 | 6,595 |
| Age | 17.4 | 18.0 | 21.0 | 25.7 | 23.6 | 25.3 |
| Black | 0.85 | 0.91 | 0.86 | 0.96 | 1.00 | 0.95 |
| Male | 0.90 | 0.98 | 0.92 | 1.00 | 1.00 | 0.90 |
| Prior Arrests | 2.5 | 8.8 | 9.2 | 16.7 | 20.3 | 13.3 |
| Share Co-Arrests | 0.21 | 0.63 | 0.64 | 0.32 | 0.45 | 0.55 |
| Degree | 2.23 | 9.56 | 9.02 | 5.02 | 9.18 | 6.38 |







Other networks



Defining potential outcomes & treatment effects

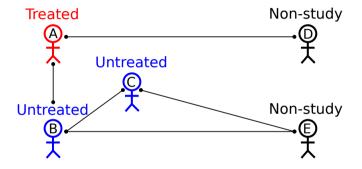
- ► If in original RCTs: 4 potential outcomes for 4 exposure states (ℰ)
 - $Y_{1,1}$: Directly treated, peer exposure $Y_{1,0}$: Directly treated, no peer exposure $Y_{0,1}$: Control, peer exposure $Y_{0,0}$: Control, no peer exposure
- ► Assume equal exposure effect for T and C (based on data and to help power)
 - $E(Y_{1,1}) E(Y_{1,0}) = E(Y_{0,1}) E(Y_{0,0})$

$$Exposure_{inRCT} = E(Y_{.,1}) - E(Y_{.,0})$$

▶ **If not in original RCTs**: 2 potential outcomes for 2 exposure states Y_1 : Peer exposure Y_0 : No peer exposure

$$Exposure_{outRCT} = E(Y_1) - E(Y_0)$$

Estimation: Aronow & Samii 2017

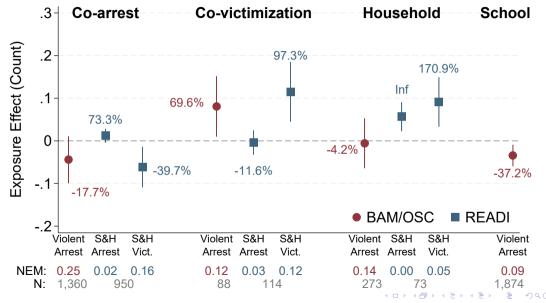


▶ Can calculate exact $p_{is}(j,k)$ or $p_{is}(k) = Pr(\mathscr{E} = e)$ for every exposure state

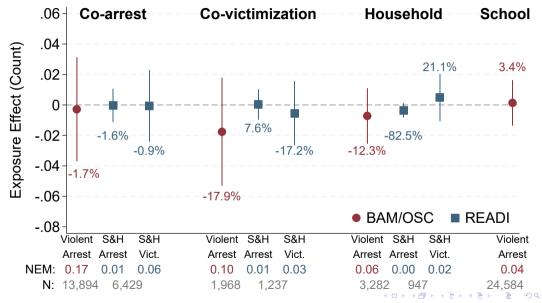
Estimation: Aronow & Samii 2017

- ► IPW: $E(\tilde{Y}_e) = \frac{1}{N} \sum_i \frac{\mathbb{1}(\mathscr{E} = e)Y_i}{p_i \cdot (\mathscr{E} = e)} \forall i \text{ where } p_{is}(\mathscr{E} = e) > 0$
- ► Estimate exposure effect= $E(\tilde{Y}_{..1}) E(\tilde{Y}_{..0})$ or $E(\tilde{Y}_1) E(\tilde{Y}_0)$
 - ▶ We do not constrain exposure effect to be the same in- and out-of-RCTs
 - ightharpoonup Out-of-RCT peers look very different \rightarrow allows for heterogeneity
 - And solves estimation problem $\rightarrow Pr(\mathscr{E} = j, k) = 0$ for out-of-RCT sample
- Residualize covariates (Aronow & Samii 2017, Särndal et al. 1992)
 - Assume $E(Y_{1,1}) E(Y_{1,0}) = E(Y_{0,1}) E(Y_{0,0}) \rightarrow \text{residualize direct treatment}$
- A & S derive asymptotic standard errors accounting for interdependence & covariates, estimates are conservative
 - Assumes exposure states are correctly specified

Violence spillovers, within RCT sample



Violence spillovers, non-RCT sample



What could explain the heterogeneity?

- ► In-RCT sample observably different in at least 3 ways:
 - lacktriangle Higher risk of violence ightarrow non-exposed means 2-3 times those for non-RCT peers
 - lacktriangle More exposures ightarrow 26 vs 11% have multiple exposures, avg 1.6 vs 1.3 | exposure
 - ightharpoonup Peers with higher take-up ightarrow 25-50% more likely to have peer who participated
- ► In-RCT sample unobservably different
 - RCT sample selection may succeed in identifying those responsive to change
 - Could use differences in selection across RCTs to assess

How big are these effects?

- ▶ Using estimates with $p \le 0.1$, assume exposure effects are additive & re-calculate social impact of RCTs
- ▶ BAM/OSC: 48 RCT members indirectly exposed (in-study 1st degree peer in any network) per 100 treated people
 - ► Stacked ITT estimate: -2 violent-crime arrests per 100 treated (-22%)
 - ▶ Indirect exposure, all networks = -0.8 arrests per 100 treated (-20%)
 - ► Accounting for spillovers, average decline in violence from RCTs 40% higher
- ▶ READI: 63 exposed per 100 treated via co-arrest, co-victim, household
 - ▶ Net RCT impact shifts from -2.2 to -2.0 per 100 T for shooting & homicide arrests
 - ► From -1.3 (insig) to -4.0 per 100 T for shooting & homicide victimization
 - ► Together ~80% larger shooting & homicide decline

Summary and next steps

- Networks dense enough to really matter
 - ► Current estimates: original RCTs may understate net violence decline by 40-80%
- Spillovers seem limited to those in original studies
 - ▶ Maybe programs successfully targeted those at risk of violence + responsive?
- Type of social relationship seems to matter
 - lacktriangle Violence \downarrow diffuses via denser networks, weaker ties. Closer ties ightarrow viol. substitutes
 - Potentially important lessons about joint crime decisions here
- Much work left to do
 - ► Geography, other exposure definitions, ties via multiple networks, school outcomes
 - Heterogeneity to inform why peer behavior matters
 - ▶ Use results, treatment effects on future tie formation to build model
 - Optimal targeting

Appendix Slides

Appendix

Household network

- Observe address and parent/guardian in CPS data once/year while a student
 - ▶ 96% of RCT members linked to non-missing info
- ► At any point in 2008/09 data (back to 90s) randomization date, link anyone who shares:
 - ► Parent/guardian name, address, and last name
 - ▶ Drop co-habitants < age 12 at randomization
 - ▶ ~90% linked peers are ages 6-25, some older co-residents
- ▶ Among RCT members in CPS data, 31% have sibling/co-habitant

Shared class network

- ▶ Use clustering algorithm (modularity blocking + stochastic block model) to identify "academic communities"
 - In year prior to randomization at school with most days present
- - Students linked if in shared academic community, not if otherwise
- Have experimented with threshold definitions (at least X shared courses)
- Schedule data currently available for OSC 1 & 2 in regular public schools (BAM coming, READI school involvement often old)

What academic communities look like

| | Mean | SD | $10 \mathrm{th}$ | $90 \mathrm{th}$ |
|------------------------|------|-----|------------------|------------------|
| Community size | 13.7 | 8.3 | 6 | 24 |
| Total classes enrolled | 15.7 | | | |
| Classes shared | | | | |
| within community | 4.8 | 3.4 | 1.4 | 9.6 |
| outside community | 0.5 | 0.4 | 0.2 | 0.9 |

Community-level observations, N = 2,072.

Class network: stochastic block model

- Number of shared classes c between 2 students i, j function of unobserved academic network A
- ▶ Want to use data to identify peers in same academic network
- ► Can write down joint distribution of network membership and avg classes shared, where q is number of academic communities
- If we knew A, could estimate α, λ via maximum likelihood
 - Instead, need to add some structure with distributional functional forms (factorized multinomial)
 - ▶ Then jointly estimate community membership and distributional parameters

Class network: stochastic block model

- ▶ Too complex to do for full data: restrict size of networks within schools
- \triangleright 2/3 students take at least 2/3 of courses with students in other grades
 - So grade blocking is limiting
- ► Use modularity clustering algorithm very fast (Clauset et al 2004)
 - ► Finding groups that maximize the difference between actual ties and what ties would look like if random
- ightharpoonup Right now, 5% of communities have <1 shared classes with each other
 - Will likely eventually trim some "weak" communities

Non-RCT estimation in regression form

1. Control for p_{is} directly, propensity-score style:

$$Y_{is} = \tau Exposed_{is} + \delta_{s}p_{is} + \gamma_{s} + \theta X_{is} + \varepsilon_{is}$$

- 2. Borusyak & Hull 2021 improvement: Recenter indirect treatment by using $Exposed_{is} - p_{is} + randomizaton inference$
 - ▶ BUT: when 2 kinds of treatments (direct, indirect), correct inference is on other side of frontier (not a sharp null)
 - ▶ Athey et al (2018): Fix direct T for subset & re-randomize for others large power reductions + estimates can vary depending on whose direct T is fixed

Co-victim first-degree peers

| | BAM/OSC | | | READI | | |
|-----------------------|---------|------------------------------------|-------------------------------------|--------|------------------------------------|-------------------------------------|
| | In-RCT | In-RCT + Potentially Exposed | Non-RCT + Potentially Exposed | In-RCT | In-RCT + Potentially Exposed | Non-RCT + Potentially Exposed |
| N | 9,590 | 88 | 1,968 | 2,456 | 114 | 1,237 |
| Age | 17.4 | 17.5 | 24.4 | 25.7 | 23.9 | 29.9 |
| Black | 0.85 | 0.98 | 0.93 | 0.96 | 1.00 | 0.91 |
| Male | 0.90 | 0.89 | 0.69 | 1.00 | 1.00 | 0.69 |
| Prior Arrests | 2.5 | 3.8 | 2.8 | 16.7 | 21.9 | 5.8 |
| Prior Victimizations | 0.6 | 2.2 | 3.2 | 1.5 | 2.3 | 3.4 |
| Degree | 0.22 | 1.85 | 2.55 | 0.46 | 2.25 | 2.10 |

Household first-degree peers

| | $_{ m BAM/OSC}$ | | | READI | | |
|---------------|-----------------|------------------------------------|-------------------------------------|--------|------------------------------------|-------------------------------------|
| | In-RCT | In-RCT + Potentially Exposed | Non-RCT + Potentially Exposed | In-RCT | In-RCT + Potentially Exposed | Non-RCT + Potentially Exposed |
| N | 9,590 | 273 | 3,847 | 2,456 | 73 | 947 |
| Age | 17.4 | 17.0 | 16.9 | 25.7 | 23.7 | 22.8 |
| Black | 0.85 | 0.88 | 0.82 | 0.96 | 0.99 | 0.98 |
| Male | 0.90 | 0.86 | 0.50 | 1.00 | 1.00 | 0.50 |
| Prior Arrests | 2.5 | 1.9 | 1.3 | 16.7 | 16.3 | 3.8 |
| Degree | 0.44 | 1.62 | 1.90 | 0.43 | 1.82 | 2.00 |

Classroom first-degree peers

| | OSC1/OSC2 | | | | |
|-----------------|-----------|------------------------------------|-------------------------------------|--|--|
| | In-RCT | In-RCT + Potentially Exposed | Non-RCT + Potentially Exposed | | |
| N | 6,850 | 1,923 | 25,202 | | |
| Age | 18.1 | 17.3 | 17.0 | | |
| Black | 0.92 | 0.96 | 0.74 | | |
| \mathbf{Male} | 0.85 | 0.65 | 0.49 | | |
| Prior Arrests | 3.2 | 1.6 | 0.7 | | |
| Degree | 6.65 | 15.32 | 18.27 | | |